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A NOTE ON CONFIDENCE BOUNDS CONNECTED WITH THE HYPOTHESIS
OF EQUALITY OF TWO DISPERSION MATRICES

by

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1. Summary and introduction. In some previous work by the author and R. Gnanadesikan [1,2], among other results, confidence bounds with a confidence coefficient greater than or equal to $1 - \alpha$, on the characteristic roots $c(\Sigma_1 \Sigma_2^{-1})$'s of Σ_1 and Σ_2 from $N(\underline{x}_1, \Sigma_1)$ and $N(\underline{x}_2, \Sigma_2)$ were given in the form (numbered (2.6) in [1])

$$(1.1) \quad \lambda_1 c_{\max}(S_1 S_2^{-1}) \geq \text{all } c(\Sigma_1 \Sigma_2^{-1}) \geq \lambda_2 c_{\min}(S_1 S_2^{-1}),$$

where (for $i = 1, 2$) S_i denotes the dispersion matrix of a random sample of size $n_i + 1$ from $N(\underline{x}_i, \Sigma_i)$, c_{\max} and c_{\min} denote the largest and smallest characteristic roots of $S_1 S_2^{-1}$ and λ_1 and λ_2 are two constants depending upon α , n_1 , n_2 , p (the number of variates) and the (known) joint distribution of the smallest and largest roots of $S_1 S_2^{-1}$ under $H_0: \Sigma_1 = \Sigma_2$.

Furthermore, with any one variate i ($i = 1, 2, \dots, p$) cut out and the (truncated) sample and population dispersion matrices being denoted by $S_1^{(i)}$, $S_2^{(i)}$, $\Sigma_1^{(i)}$, $\Sigma_2^{(i)}$, or with any two variates i and j ($i \neq j = 1, 2, \dots, p$) cut out and the (truncated) sample and population dispersion matrices being denoted by $S_1^{(i,j)}$, $S_2^{(i,j)}$, $\Sigma_1^{(i,j)}$ and $\Sigma_2^{(i,j)}$, and so on, the following types of confidence bounds (numbered (2.7) and (2.8) in [1]) were given:

$$(1.2) \quad \lambda_1 c_{\max}(S_1^{(i)} S_2^{(i)-1}) \geq \text{all } c(\Sigma_1^{(i)} \Sigma_2^{(i)-1}) \geq \lambda_2 c_{\min}(S_1^{(i)} S_2^{(i)-1})$$

and

$$(1.3) \quad \lambda_1 c_{\max} (S_1^{(i,j)} S_2^{(i,j)^{-1}}) \geq \text{all } c(\Sigma_1^{(i,j)} \Sigma_2^{(i,j)^{-1}}) \\ \geq \lambda_2 c_{\min} (S_1^{(i,j)} S_2^{(i,j)^{-1}}),$$

with obvious extensions to truncations of a higher order. However, in obtaining the truncated confidence bounds like (1.2) and (1.3) (numbered (2.7) and (2.8) in [1]) from (2.3) of [1] certain steps were omitted, thus rendering the derivation abrupt. It is the purpose of the present note to supply the missing steps and make the derivation complete.

2. Derivation of the confidence bounds on the roots connected with the truncated matrices. We go back to [1] and note that it is perfectly clear how the passage from (2.1) to (2.3) and from (2.3) on to (2.6), which is the confidence statement on the total matrices, are made. Now starting from (2.1) or rather from the step just back of (2.1), we shall deduce a result on truncated matrices which is exactly analogous to and would be implied by (2.1) on the total matrices. The passage from this analogous result to the analogous of (2.3) and from there on to (2.7) and (2.8) of [1] will obviously be exactly identical with the passage from (2.1) to (2.3) and from there on to (2.6). Back in (2.1) of [1], we observe that the step back of (2.1) of [1] is the distribution of $Y_1(p \times n_1)$ and $Y_2(p \times n_2)$ given by

$$(2.1) \quad \text{Const. exp } \int -\frac{1}{2} \text{tr} (\Sigma_1^{-1} S_1 + \Sigma_2^{-1} S_2) \int dY_1 dY_2,$$

where $Y_1 Y_1' = n_1 S_1$ and $Y_2 Y_2' = n_2 S_2$, and where S_1, S_2, Σ_1 and Σ_2 have been already defined. Put $\Sigma_1 = \tau_1 \tau_1'$ and $\Sigma_2 = \tau_2 \tau_2'$, where τ_1 and τ_2 are lower

triangular matrices and notice that

$$\begin{aligned} \text{tr} (\Sigma_1^{-1} S_1 + \Sigma_2^{-1} S_2) &= \text{tr} \left[(\tau_1 \tau_1')^{-1} S_1 + (\tau_2 \tau_2')^{-1} S_2 \right] \\ &= \text{tr} \left[\tau_1^{-1} S_1 \tau_1'^{-1} + \tau_2^{-1} S_2 \tau_2'^{-1} \right]. \end{aligned}$$

As before, we note that the distribution of $c(S_1 S_2^{-1})$'s when $\Sigma_1 = \Sigma_2$ is exactly the same as of $c[\tau_1^{-1} S_1 \tau_1'^{-1} (\tau_2^{-1} S_2 \tau_2'^{-1})^{-1}]$'s when $\Sigma_1 \neq \Sigma_2$.

Thus given α , we can find λ_1 and λ_2 and make, with a probability $1 - \alpha$, the statement

$$(2.2) \quad \lambda_1 \geq \text{all } c \left[(\tau_2^{-1} S_2 \tau_2'^{-1}) (\tau_1^{-1} S_1 \tau_1'^{-1}) \right] \geq \lambda_2,$$

which is exactly equivalent to (2.1) of $[1]$ and also to

$$(2.3) \quad \lambda_1 \geq \frac{\underline{a}' \left[\tau_2^{-1} S_2 \tau_2'^{-1} \right] \underline{a}}{\underline{a}' \left[\tau_1^{-1} S_1 \tau_1'^{-1} \right] \underline{a}} \geq \lambda_2,$$

for all nonnull \underline{a} . Let us now partition τ_1 and τ_2 into

$$(2.4) \quad \tau_1 = \begin{bmatrix} \tau_{11} & 0 \\ \tau_{12} & \tau_{13} \end{bmatrix} \begin{matrix} r \\ p-r \end{matrix}, \quad \tau_2 = \begin{bmatrix} \tau_{21} & 0 \\ \tau_{22} & \tau_{23} \end{bmatrix} \begin{matrix} r \\ p-r \end{matrix},$$

where, of course, τ_{11} , τ_{13} , τ_{21} , τ_{23} are lower triangular, and τ_{12} and τ_{22} are solid. A so put

$$(2.5) \quad S_1 = T_1 T_1', \quad S_2 = T_2 T_2',$$

where T_1 and T_2 are lower triangular matrices, and partition T_1 and T_2 into

$$(2.6) \quad T_1 = \begin{array}{ccc} T_{11} & 0 & r \\ T_{12} & T_{13} & p-r \\ r & p-r & \end{array}, \quad T_2 = \begin{array}{ccc} T_{21} & 0 & r \\ T_{22} & T_{23} & p-r \\ r & p-r & \end{array}$$

where, as in (2.4), T_{11} , T_{13} , T_{21} , T_{23} are lower triangular, and T_{12} and T_{22} are solid. It is easy to check that τ_1^{-1} and τ_2^{-1} are also lower triangular matrices and can be expressed as

$$(2.7) \quad \tau_1^{-1} = \begin{array}{ccc} \tau_{11}^{-1} & 0 & r \\ -\tau_{13}^{-1} \tau_{12} \tau_{11}^{-1} & \tau_{13}^{-1} & p-r \\ r & p-r & \end{array} \quad \text{and}$$

$$\tau_2^{-1} = \begin{array}{ccc} \tau_{21}^{-1} & 0 & r \\ -\tau_{23}^{-1} \tau_{22} \tau_{21}^{-1} & \tau_{23}^{-1} & p-r \\ r & p-r & \end{array} .$$

Using (2.5), (2.6) and (2.7) we can now rewrite (2.3) as

$$(2.8) \quad \lambda_1 \geq \frac{\underline{a}' \left[\tau_2^{-1} T_2 T_2' \tau_2^{-1} \right] \underline{a}}{\underline{a}' \left[\tau_1^{-1} T_1 T_1' \tau_1^{-1} \right] \underline{a}} \geq \lambda_2 ,$$

for all nonnull \underline{a} , where

$$(2.9) \quad \tau_2^{-1} T_2 = \begin{array}{ccc} \tau_{21}^{-1} T_{21} & 0 & r \\ -\tau_{23}^{-1} \tau_{22} \tau_{23}^{-1} T_{21} + \tau_{23}^{-1} T_{22} & \tau_{23}^{-1} T_{23} & p-r \\ r & p-r & \end{array}$$

and

$$(2.10) \quad \tau_1^{-1} T_1 = \begin{bmatrix} \tau_{11}^{-1} T_{11} & 0 \\ -\tau_{13}^{-1} \tau_{12} \tau_{13}^{-1} T_{11} + \tau_{13}^{-1} T_{12} & \tau_{13}^{-1} T_{13} \end{bmatrix} \begin{matrix} r \\ p-r \end{matrix}$$

The matrices $T_2' \tau_2'^{-1}$ and $T_1' \tau_1'^{-1}$ are the transposes of the right sides of (2.9) and (2.10) and can be easily written in the analogous partitioned forms. If we now specialize \underline{a} so that the last $p-r$ components are zero, and denote by \underline{a}^* the r -dimensional vector with the first r components of the specialized \underline{a} , then it is easy to check that (2.8) \implies

$$(2.11) \quad \lambda_1 \geq \frac{\underline{a}^{*'} \int \tau_{21}^{-1} T_{21} T_{21}' \tau_{21}'^{-1} \int \underline{a}^*}{\underline{a}^{*'} \int \tau_{11}^{-1} T_{11} T_{11}' \tau_{11}'^{-1} \int \underline{a}^*} \geq \lambda_2,$$

for all nonnull (r -dimensional) \underline{a}^* . Furthermore, notice that

$$(2.12) \quad \begin{aligned} S_1^{(r+1, \dots, p)} &= T_{11} T_{11}', & S_2^{(r+1, \dots, p)} &= T_{21} T_{21}', \\ \Sigma_1^{(r+1, \dots, p)} &= \tau_{11} \tau_{11}', & \Sigma_2^{(r+1, \dots, p)} &= \tau_{21} \tau_{21}', \end{aligned}$$

where the symbols on the left side of all the equations stand for the truncated matrices obtained by cutting out the last $p-r$ variates. Now since (2.3) \iff (2.2) \iff (2.1) of $\int 1 \int \iff$ (2.3) of $\int 1 \int \iff$ (2.6) of $\int 1 \int$, therefore, (2.11) \iff (2.13) analogous to (2.2), but on the truncated matrices \iff (2.14) analogous to (2.1) of $\int 1 \int$, but on the truncated matrices \iff (2.15) analogous to (2.3) of $\int 1 \int$, but on the truncated matrices \implies (2.16) analogous to (2.6) of $\int 1 \int$, but on the truncated matrices. Now, since (2.8) \implies (2.11) and (2.11) \implies (2.16),

we have, with a confidence coefficient $\geq 1-\alpha$, the confidence statement.

$$(2.16) \quad \lambda_1 c_{\max} \left[S_1^{(r+1, \dots, p)} S_2^{(r+1, \dots, p)^{-1}} \right] \\ \geq \text{all } c \left[\Sigma_1^{(r+1, \dots, p)} \Sigma_2^{(r+1, \dots, p)^{-1}} \right] \\ \geq \lambda_2 c_{\min} \left[S_1^{(r+1, \dots, p)} S_2^{(r+1, \dots, p)^{-1}} \right].$$

It is clear that, starting from (2.1), we could have cut out any $p-r$ variates instead of the last $(p-r)$ ones (with $r = 1, 2, \dots, p-1$), and thus we have, with a confidence coefficient $\geq 1-\alpha$, the set of $\binom{p}{2}$ confidence statements of the type (1.2) (with $i = 1, 2, \dots, p$), of the type (1.3) (with $i \neq j = 1, 2, \dots, p$), and so on.

References

- [1] Roy, S. N. and Gnanadesikan, R., "Further contributions to multi-variate confidence bounds," North Carolina Institute of Statistics Mimeograph Series No. 155.
- [2] Roy, S. N. and Gnanadesikan, R., "Further contributions to multi-variate confidence bounds," Biometrika, Vol. 44 (1957), pp. 399-410.