

SOME RESULTS USING GENERAL MOMENT FUNCTIONS

by

Walter L. Smith

University of North Carolina

Institute of Statistics Mimeo Series No. 554

This research was supported by the Department of the  
Navy, Office of Naval Research. Grant NONR-855(09).

DEPARTMENT OF STATISTICS

University of North Carolina

Chapel Hill, N. C.

## 1. Introduction

The work of the present note originated in a desire to prove Theorem 1 below, which, in a somewhat more restricted version, was announced by the author during the Chapel Hill Symposium on Congestion Theory (Smith and Wilkinson (1965), pp. 132-133). The proof then available was needlessly complicated; it will now follow from a useful general theorem on generating functions (Corollary 3.1, below).

In Smith (1967) a class  $\mathcal{M}$  of "moment functions"  $M(x)$  is introduced as a means of dealing with problems involving a random variable  $X$ , say, when it is given that the highest absolute moment of  $X$  known to be finite is not necessarily an integral power of  $X$ . For example, one might be given that  $\mathcal{E}|X|^{3/2} \log^+|X| < \infty$ . In Smith (1967) the use of such information in the expansion of characteristic functions and in certain renewal-theoretic problems is discussed. A rather more fundamental property of a large class of these functions  $M$  is as follows. If  $X_1, X_2, \dots$ , are independent and identically distributed random variables such that  $0 < \mu = \mathcal{E}X_n < \infty$  and if  $S_n = X_1 + \dots + X_n$ , then provided  $\mathcal{E}M(|X|) < \infty$  it follows that  $\mathcal{E}M(|S_n|) \sim M(n\mu)$  as  $n \rightarrow \infty$ . This result is described in detail in Theorem 2.

Theorem 3, below, is then an easy consequence of Theorem 2. It concerns the finiteness of  $\mathcal{E}M(|S_N|)$  when  $N$  is a non-negative integer-valued random variable independent of the  $\{X_n\}$ . When the  $\{X_n\}$  are themselves also non-negative integer-valued random variables we are led at once to Corollary 3.1, below, concerning a useful closure property of

probability generating functions of probability generating functions. As explained, it is from this Corollary that we deduce Theorem 1, the original object of the present work.

We now state:

Theorem 1 (i) Let  $\{X_n\}$  be an infinite sequence of independent and identically distributed random variables, with common distribution function

$$F(x) = P\{X_n \leq x\}, \text{ such that } 0 < \int_{-\infty}^{\infty} x dF(x) < +\infty.$$

(ii) Write  $S_n = X_1 + \dots + X_n$ , as usual, and  $T_n = \max\{S_1, S_2, \dots, S_n\}$ .

(iii) Let  $\nu \geq 0$  be an integer and  $M(x)$  be a non-decreasing function

of  $x \geq 0$  such that  $M(x)/x$  is non-increasing for all large  $x$ . Suppose

$M(0) > 0$  (and note that  $M(x)$  may be a constant identically). Then any

one of the following three statements implies the other two:

$$(1.1) \quad \int_{-\infty}^0 |x|^{(\nu+1)} M(|x|) dF(x) < \infty;$$

$$(1.2) \quad \sum_{n=1}^{\infty} n^{(\nu-1)} M(n) P\{S_n \leq 0\} < \infty;$$

$$(1.3) \quad \sum_{n=1}^{\infty} n^{\nu} M(n) P\{T_n \leq 0\} < \infty.$$

In order to prove this result on random walks, which, incidentally is of some relevance to the theory of queues, we prove the following Theorem 2. This theorem would appear to be of general interest.

Theorem 2 Suppose the conditions and notations (i), (ii) and (iii) of Theorem 1 hold, with the additional proviso that  $\mu = \int_{-\infty}^{\infty} x dF(x) < \infty$ .

Then, if  $\mathcal{E}|X_1|^\nu M(|X_1|) < \infty$ , it follows that, as  $n \rightarrow \infty$ ,

$$\mathcal{E}|S_n|^\nu M(|S_n|) \sim (n\mu)^\nu M(n\mu).$$

If, however,  $\mu = \mathcal{E}X_n = \infty$  (which can be the case only when  $\nu = 0$ )

one can infer from  $\mathcal{E}M(|X_1|) < \infty$  only the weaker conclusion  $\mathcal{E}M(|S_n|) = o(n)$ .

From Theorem 2, as already explained, we can easily deduce the following theorem concerning moments of sums of random numbers of random variables.

Theorem 3 (i) Let  $\{X_n\}$  be an infinite sequence of independent and identi-

cally distributed random variables such that  $0 < \mathcal{E}X_n < \infty$  and write

$$\underline{B}_n(x) = P\{S_n \leq x\} \text{ for } n = 1, 2, \dots; \underline{B}_0(x) = P\{0 \leq x\}.$$

(ii) Let  $\{c_n\}$  be a discrete probability distribution (i.e.  $c_n \geq 0$ ,

all  $n$ , and  $\sum_0^\infty c_n = 1$ ) such that  $c_0 < 1$ .

(iii) Let the d.f.  $A(x)$  be thus defined:

$$A(x) = \sum_{n=0}^{\infty} c_n \underline{B}_n(x)$$

(iv) Let  $\nu$  and  $M(x)$  be as in Theorem 1 (iii).

Then a necessary and sufficient condition to ensure that

$$(1.4) \quad \int_{-\infty}^{+\infty} |x|^\nu M(|x|) dA(x) < \infty$$

is that both the following statements be true:

$$(1.5) \quad \int_{-\infty}^{+\infty} |x|^{\nu} M(|x|) dB_1(x) < \infty;$$

$$(1.6) \quad \sum_{n=1}^{\infty} n^{\nu} M(n) c_n < \infty.$$

If, for the case  $\nu = 0$ , it should happen that (i), ..., (iv) hold except for the fact that  $\sum_{n=1}^{\infty} c_n = +\infty$  then one can still infer that (1.4) is true if (1.5) is given, together with the condition  $\sum_{n=1}^{\infty} n c_n < \infty$ .

If  $\nu$  and  $M(x)$  are as in Theorem 1 and if  $\Pi(s) = \sum_{n=0}^{\infty} \pi_n s^n$ ,  $0 \leq s \leq 1$ , is a probability generating function, then we shall say that  $\Pi$  is in the class  $\mathcal{G}(\nu, M)$  if

$$\sum_{n=1}^{\infty} n^{\nu} M(n) \pi_n < \infty.$$

An immediate corollary to Theorem 3 is the following:

Corollary 3.1 Let  $\nu$  and  $M(x)$  be as in Theorem 1 (iii) and let  $\alpha(s)$ ,  $\beta(s)$ ,  $\gamma(s)$  be probability generating functions,  $0 \leq s \leq 1$ . Suppose that  $\beta(s) = \sum_{n=0}^{\infty} b_n s^n$ , that  $\sum_{n=0}^{\infty} n b_n < \infty$ , and that  $\alpha(s) = \gamma(\beta(s))$ . Then  $\alpha \in \mathcal{G}(\nu, M)$  if and only if both  $\beta$  and  $\gamma$  belong to  $\mathcal{G}(\nu, M)$ .

If, for the case  $\nu = 0$ , it should happen that  $\sum_{n=1}^{\infty} n b_n = \infty$  one can still infer that  $\alpha \in \mathcal{G}(0, M)$  if it is given that  $\beta \in \mathcal{G}(0, M)$  and  $\sum_{n=1}^{\infty} n c_n < \infty$ .

We close this introduction by mentioning that in the course of proving Theorem 2 we establish the following:

Corollary 2.1 Under the conditions of Theorem 1, for any  $\epsilon > 0$ ,

$$\int_{\left| \frac{S_n}{n} - \mu \right| > \epsilon} |S_n|^{\nu} M_{\nu}(|S_n|) dP = o(n^{\nu} M(n)),$$

as  $n \rightarrow \infty$ .

2. On moment functions and some preliminary lemmas.

In Smith (1967) we have introduced a class  $\mathfrak{M}$  of moment-functions. A function  $M(x)$  belongs to  $\mathfrak{M}$  if: (i) it is non-decreasing in  $[0, \infty)$ ; (ii)  $M(x) \geq 1$ , all  $x \geq 0$ ; (iii)  $M(x+y) \leq M(x)M(y)$  for all  $x, y \geq 0$ . If  $M \in \mathfrak{M}$  and, additionally, (iv)  $M(2x) = O(M(x))$  for all  $x \geq 0$  then we say  $M \in \mathfrak{M}^*$ . We say a d.f.  $F(x)$  is in the class  $\mathcal{D}(\nu, M)$ , for any real  $\nu \geq 0$ , if  $\int_{-\infty}^{+\infty} |x|^\nu M(|x|) dF(x) < \infty$ . From the point of view of defining a class  $\mathcal{D}(\nu, M)$ , the moment function  $M(x)$  can be replaced by any non-decreasing function  $N(x)$  such that  $N(x) \asymp M(x)$  as  $x \rightarrow \infty$  (" $\phi(x) \asymp \psi(x)$  as  $x \rightarrow \infty$ " means that there exist positive constants  $\delta$  and  $\Delta$  such that  $\delta\phi(x) \leq \psi(x) \leq \Delta\psi(x)$  for all large  $x$ ). Thus we are led to introduce an extended class  $E\mathfrak{M}^*$  in such a way that a function  $N(x) \in E\mathfrak{M}^*$  if we can find a function  $M(x) \in \mathfrak{M}^*$  such that  $N(x) \asymp M(x)$  as  $x \rightarrow \infty$ . We can then say  $M(x)$  represents  $N(x)$  in  $\mathfrak{M}^*$ .

Suppose that for some  $\Delta \geq 0$  and some  $A > 0$  the function  $N(x)$  satisfies the following:

- (i)  $N(x)$  is non-decreasing for  $x \geq \Delta$ ;
- (ii)  $N(\Delta) > 0$ ;
- (iii)  $N(x+y) \leq AN(x)N(y)$  for all  $x \geq \Delta, y \geq \Delta$ ;
- (iv)  $N(2x) = O(N(x))$  for all  $x \geq \Delta$ .

Then we may suppose, with no loss of generality, that  $AN(\Delta) \geq 1$  and define

$$\begin{aligned} M(x) &= AN(x), & x \geq \Delta, \\ &= M(\Delta), & x \leq \Delta; \end{aligned}$$

the function  $M(x)$  so defined belongs to  $\mathcal{M}^*$ . Thus a function  $N(x)$  satisfying (i), (ii), (iii) and (iv) above belongs to  $E\mathcal{M}^*$ .

Now suppose that, for some  $\Delta \geq 2$ ,  $M(x)$  is non-decreasing and  $M(x)/x$  is non-increasing for all  $x \geq \Delta$ , and  $M(\Delta) > 0$ . Let  $\nu \geq 0$  be a constant, and define  $N_\nu(x) = x^\nu M(x)$  for all  $x \geq \Delta$ . It is clear that (i) and (ii) above are satisfied by  $N_\nu(x)$ . Further, if  $x \geq \Delta$  and  $y \geq \Delta$ ,

$$\frac{N_\nu(x+y)}{N_\nu(x)N_\nu(y)} = \frac{(x+y)^\nu}{x^\nu y^\nu} \cdot \frac{(x+y)}{M(x)M(y)} \left\{ \frac{M(x+y)}{x+y} \right\}$$

But, since  $M(x)/x$  is non-increasing,

$$(x+y) \left\{ \frac{M(x+y)}{x+y} \right\} \leq x \left\{ \frac{M(x)}{x} \right\} + y \left\{ \frac{M(y)}{y} \right\},$$

so that

$$\begin{aligned} \frac{N_\nu(x+y)}{N_\nu(x)N_\nu(y)} &\leq \left( \frac{1}{x} + \frac{1}{y} \right)^\nu \left\{ \frac{1}{M(x)} + \frac{1}{M(y)} \right\} \\ &\leq \frac{2}{M(\Delta)}, \end{aligned}$$

since  $x \geq \Delta$ ,  $y \geq \Delta$ . Thus  $N(x)$  satisfies (iii) above. Finally, for  $x \geq \Delta$ ,

$$\begin{aligned} \frac{N_\nu(2x)}{N_\nu(x)} &\leq 2^\nu \frac{M(2x)}{M(x)} \\ &= 2^{(\nu+1)} \frac{x}{M(x)} \left\{ \frac{M(2x)}{2x} \right\} \\ &\leq 2^{(\nu+1)} \frac{x}{M(x)} \left\{ \frac{M(x)}{x} \right\} \end{aligned}$$

Thus  $N_\nu(x)$  also satisfies (iv) and so belongs to  $\mathcal{E}\mathcal{M}^*$  and there will be a function  $M_\nu(x) \in \mathcal{M}^*$  which represents  $N_\nu(x)$ . We shall make frequent use of these functions  $M_\nu(x)$ .

With these preliminaries set aside we can now prove the following.

Lemma 2.1 Let  $k \geq 1$  be an integer and let  $M(x) \in \mathcal{M}$ . Suppose  $X_1, X_2, \dots, X_k$  are independent and identically distributed random variables and write  $S_j = X_1 + \dots + X_j$  for  $j = 1, 2, \dots, k$ . Then  $\mathcal{E}M(|S_k|) < \infty$  if and only if  $\mathcal{E}M(|X_1|) < \infty$ .

Proof. Clearly  $M(|S_k|) \leq M(\sum_1^k |X_j|)$ , and so, by a property of functions in  $\mathcal{M}$ ,

$$M(|S_k|) \leq \prod_{j=1}^k M(|X_j|).$$

Therefore

$$\mathcal{E}M(|S_k|) \leq [\mathcal{E}M(|X_1|)]^k,$$

and part of the lemma is proved.

If  $P\{X_1 \geq 0\} = 1$ , or if  $P\{X_1 \leq 0\} = 1$  the rest of the proof is trivial. Suppose therefore that both  $P\{X_1 \geq 0\} > 0$  and  $P\{X_1 < 0\} > 0$ . Then both  $P\{S_{k-1} \geq 0\} > 0$  and  $P\{S_{k-1} < 0\} > 0$ . The finiteness of  $\mathcal{E}M(|S_k|)$  then implies that, for example,

$$\mathcal{E}\{M(|S_k|) | S_{k-1} < 0; X_k < 0\}$$

is also finite. However, if both  $S_{k-1}$  and  $X_k$  are negative then  $|X_k| < |S_k|$  and so

$$\mathcal{E}\{M(|X_k|) | S_{k-1} < 0; X_k < 0\}$$

must be finite. But  $X_k$  and  $S_{k-1}$  are independent. Thus  $\mathcal{E}\{M(|X_k|) | X_k < 0\}$



is finite. Similarly one can deal with  $\{M(|X_k|) | X_k \geq 0\}$ , and the lemma follows.

Lemma 2.2 If, for large x, L(x) is non-increasing and xL(x) is non-decreasing, then

$$\lim_{\epsilon \rightarrow 0+} \limsup_{n \rightarrow \infty} \frac{\overline{L(1 + \epsilon n)}}{L(n)} = 1$$

and

$$\lim_{\epsilon \rightarrow 0+} \liminf_{n \rightarrow \infty} \frac{\overline{L(1 + \epsilon n)}}{L(n)} = 1$$

Proof If we set  $M(x) = xL(x)$ , then

$$\frac{\overline{M(1 + \epsilon n)}}{M(n)} \geq 1,$$

which implies that for all n

$$1 \geq \frac{\overline{L(1 + \epsilon n)}}{L(n)} \geq \frac{1}{1 + \epsilon}.$$

The lemma follows at once from these inequalities.

### 3. Proof of Theorem 2

Let us write  $M_\nu(x)$ ,  $\nu > 0$ , for any member of  $\mathcal{M}^*$  such that  $M_\nu(x) \asymp x^\nu M(x)$  as  $x \rightarrow \infty$ . Suppose that the integer  $\nu \geq 1$ . Then we can find a  $K_1 > 0$ ,  $K_2 > 0$ , such that, for all large  $x$ ,

$$(3.1) \quad \begin{aligned} M_\nu(x) &\leq K_1 x M_{\nu-1}(x) \\ &\geq K_2 x M_{\nu-1}(x) \end{aligned}$$

For large  $R > 0$  and small  $\varepsilon > 0$ , define events

$$A_n(\varepsilon) \equiv \{|S_n/n - \mu| \leq \varepsilon\}$$

$$B_n(R) \equiv \{|S_n/n - \mu| > R\}.$$

Further, for  $n = 1, 2, \dots$ , let

$$\Delta_\nu(R, n) = \int_{B_n(R)} \frac{M_\nu(|S_n|)}{M_\nu(n\mu)} dP.$$

Then

$$\begin{aligned} \Delta_\nu(R, n) &\leq \frac{K_1}{K_2 n \mu} \int_{B_n(R)} |S_n| \frac{M_{\nu-1}(|S_n|)}{M_{\nu-1}(n\mu)} dP \\ &\leq \frac{K_1}{K_2 \mu} \int_{B_n(R)} |X_n| \frac{M_{\nu-1}(|S_n|)}{M_{\nu-1}(n\mu)} dP, \end{aligned}$$

by symmetry and the inequality  $|S_n| \leq |X_1| + \dots + |X_n|$ .

$$\begin{aligned} \text{But } M_{\nu-1}(|S_n|) &\leq M_{\nu-1}(|S_{n-1}| + |X_n|) \\ &\leq M_{\nu-1}(|X_n|)M_{\nu-1}(|S_{n-1}|), \end{aligned}$$

by a defining property of  $\mathcal{M}^*$ . Hence we have

$$\begin{aligned} \Delta_{\nu}(R, n) &\leq \frac{K_1}{K_2^2 \mu} \int_{B_n(R)} M_{\nu}(|X_n|) \frac{M_{\nu-1}(|S_{n-1}|)}{M_{\nu-1}(n\mu)} dP \\ &\leq \frac{K_1}{K_2^2 \mu} \int_{|S_{n-1} - (n-1)\mu| > \frac{1}{2} nR} \\ &\quad + \frac{K_1}{K_2^2 \mu} \int_{|X_n - \mu| > \frac{1}{2} nR} \end{aligned}$$

$$(3.2) \quad = I_1(n) + I_2(n), \text{ say.}$$

Because  $X_n$  and  $S_{n-1}$  are independent we have

$$(3.3) \quad I_2(n) \leq \frac{K_1 \mathcal{E} M_{\nu-1}(|S_{n-1}|)}{K_2^2 \mu M_{\nu-1}(n\mu)} \int_{|X_n - \mu| > \frac{1}{2} nR} M_{\nu}(|X_1|) dP,$$

using the fact that  $X_1$  and  $X_n$  are identically distributed.

In a similar way we find

$$(3.4) \quad I_1(n) \leq \frac{K_1 \mathcal{E} M_{\nu}(|X_1|)}{K_2^2 \mu} \Delta_{\nu-1}(\frac{1}{2} R, n-1).$$

For integer  $\lambda \geq 0$ , let  $\mathcal{H}_{\lambda}$  be the hypothesis that the following two statements are true:

$$(i) \mathcal{E} M_{\lambda}(|S_n|) = O(M_{\lambda}(n\mu)), \text{ as } n \rightarrow \infty;$$

(ii) For all sufficiently large  $R$ ,  $\Delta_\lambda(R, n) \rightarrow 0$ , as  $n \rightarrow \infty$ .

The inequalities (3.2), (3.3) and (3.4) show that if  $\mathcal{E}M_\nu(|X_1|)$  is finite and if  $\mathcal{H}_{(\nu-1)}$  is true, then

$$(3.5) \quad \Delta_\nu(R, n) \rightarrow 0, \quad \text{as } n \rightarrow \infty,$$

for all large  $R$ . Furthermore,

$$(3.6) \quad \frac{\mathcal{E}M_\nu(|S_n|)}{M_\nu(n\mu)} \leq \frac{M_\nu(\overline{R + \mu n})}{M_\nu(\mu n)} + \Delta_\nu(R, n).$$

But, for some  $K_4 > 0$ ,

$$(3.7) \quad \begin{aligned} \frac{M_\nu(\overline{R + \mu n})}{M_\nu(\mu n)} &\leq \frac{K_4(R + \mu)^\nu M(\overline{R + \mu n})}{\mu^\nu M(\mu n)} \\ &\leq \frac{K_4(R + \mu)^{\nu+1}}{\mu^{\nu+1}}, \end{aligned}$$

since  $M(x)/x$  is non-increasing.

From (3.5), (3.6) and (3.7) we see that

$$\mathcal{E}M_\nu(|S_n|) = o(M_\nu(n\mu)), \quad \text{as } n \rightarrow \infty.$$

Thus  $\mathcal{H}_\lambda$  is true if  $\mathcal{H}_{(\lambda-1)}$  is true and  $M_\nu(|X_1|) < \infty$ . Plainly we desire to show that  $\mathcal{H}_0$  is true, from which it would follow (if  $\mathcal{E}M_\nu(|X_1|) < \infty$ ) that  $\mathcal{H}_\nu$  is true.

Let us write  $M(x) = xL(x)$ , where  $L(x)$  is a non-increasing function of  $x$ . Then

$$\Delta_0(R, n) = \int_{B_n(R)} \frac{|S_n| L(|S_n|)}{n\mu L(n\mu)} dP$$

$$(3.8) \quad \leq \frac{L(R - \mu n)}{\mu L(n\mu)} \int_{B_n(R)} \left| \frac{S_n}{n} \right| dP.$$

Let us define, in a familiar way, for  $n = 1, 2, \dots$ ,  $X_n^+ = X_n$  if  $X_n \geq 0$ ,  $X_n^+ = 0$  if  $X_n < 0$ . We then have  $X_n^- = X_n^+ - X_n$ ;  $\mu^+ = \mathbb{E}X_n^+$  and  $\mu^- = \mathbb{E}X_n^-$ , both necessarily finite because  $\mathbb{E}|X_n| < \infty$ ;  $S_n^+ = X_1^+ + \dots + X_n^+$ , and so on. We then have ( $R \gg \mu$ ),

$$\int_{B_n(R)} \left| \frac{S_n}{n} \right| dP = \int_{\substack{B_n(R) \\ S_n^+ \geq S_n^-}} \left| \frac{S_n}{n} \right| dP + \int_{\substack{B_n(R) \\ S_n^- > S_n^+}} \left| \frac{S_n}{n} \right| dP$$

$$(3.9) \quad \leq \int_{\substack{S_n^+ \geq n(R+\mu)}} \frac{S_n^+}{n} dP + \int_{\substack{S_n^- \geq n(R-\mu)}} \frac{S_n^-}{n} dP$$

By the weak law of large numbers, for any  $\epsilon > 0$   
 $P\{|S_n^+/n - \mu^+| \leq \epsilon\} \rightarrow 1$ , so that

$$\liminf_{n \rightarrow \infty} \int_{|S_n^+/n - \mu^+| \leq \epsilon} \frac{S_n^+}{n} dP \geq (\mu^+ - \epsilon)$$

But  $\mathbb{E}(S_n^+/n) = \mu^+$ , and hence, for  $(R + \mu) > \mu^+ + \epsilon$ ,

$$\limsup_{n \rightarrow \infty} \int_{\substack{S_n^+ \geq n(R + \mu)}} \frac{S_n^+}{n} dP \leq \epsilon.$$

The arbitrariness of  $\epsilon$  tells us that

$$\int_{S_n^+ \geq n(R+\mu)} \frac{S_n^+}{n} dP \rightarrow 0, \quad n \rightarrow \infty.$$

A similar argument will deal with the second integral in (3.9) and show that

$$\int_{B_n(R)} \left| \frac{S_n}{n} \right| dP \rightarrow 0, \quad n \rightarrow \infty.$$

This result coupled with the observation that  $L(\overline{R - \mu n}) \leq L(\mu n)$ , for  $R \gg \mu$ , enables us to deduce from (3.8) that

$$\Delta_0(R, n) \rightarrow 0, \quad n \rightarrow \infty,$$

for  $R$  sufficiently large. The proof that  $\mathcal{E}M(|S_n|) = O(M(n\mu))$  now follows easily from our inequality (3.6) with zero substituted for  $\nu$ .

Thus hypothesis  $\mathcal{H}_0$  holds and consequently so does  $\mathcal{H}_\nu$ . We shall now use the weak law of large numbers to infer that, for any  $\epsilon > 0$ ,  $P\{A_n(\epsilon)\} \rightarrow 1$  as  $n \rightarrow \infty$ . We then have

$$\begin{aligned} \int_{B_n(R) \cap \sim A_n(\epsilon)} \frac{M_\nu(|S_n|)}{M_\nu(n\mu)} &\leq \frac{M_\nu(\overline{R + \mu n})}{M_\nu(n\mu)} P\{\sim A_n(\epsilon)\} \\ &\leq K_4 \left( \frac{R + \mu}{\mu} \right)^{(\nu+1)} P\{\sim A_n(\epsilon)\} \end{aligned}$$

by an argument we have already used. Therefore, using the fact that

$\Delta_\nu(R, n) \rightarrow 0$  as  $n \rightarrow \infty$ , we can infer that

$$(3.10) \quad \mathcal{E}|S_n|^\nu M(|S_n|) = \int_{A_n(\epsilon)} |S_n|^\nu M(|S_n|) dP + o(M_\nu(n\mu)).$$

We first deduce from (3.10) that

$$(3.11) \quad \limsup_{n \rightarrow \infty} \frac{\mathcal{E}|S_n|^\nu M(|S_n|)}{(n\mu)^\nu M(n\mu)} \leq \limsup_{n \rightarrow \infty} \frac{(\mu + \epsilon)^{(\nu+1)} L(\overline{\mu + \epsilon n})}{\mu^{(\nu+1)} L(\mu n)} \\ \leq 1,$$

in view of the arbitrariness of  $\epsilon$  and Lemma 2.2. We secondly deduce from (3.10) that

$$(3.12) \quad \liminf_{n \rightarrow \infty} \frac{\mathcal{E}|S_n|^\nu M(|S_n|)}{(n\mu)^\nu M(n\mu)} \\ \geq \liminf_{n \rightarrow \infty} \frac{(\mu - \epsilon)^{(\nu+1)} L(\overline{\mu - \epsilon n})}{\mu^{(\nu+1)} L(\mu n)} P\{A_n(\epsilon)\} \\ \geq 1,$$

by the arbitrariness of  $\epsilon$ , the weak law of large numbers, and Lemma 1. The theorem follows from (3.11) and (3.12).

It will be clear from our proof that Corollary 2.1 is also established.

4. Proof of Theorem 3.

Let  $k$  be the least integer such that  $c_k > 0$ . Then it follows from the assumption  $A \in \mathcal{D}(\nu, M)$  that  $B_k \in \mathcal{D}(\nu, M)$ ; in other words, we have

$$\mathcal{E}|S_k|^\nu M(|S_k|) < \infty.$$

From Lemma 2.1 we then have that

$$\mathcal{E}|X_1|^\nu M(|X_1|) < \infty.$$

If we may suppose  $0 < \mathcal{E}X_1 < \infty$ , we can now appeal to Theorem 2 and deduce that

$$\mathcal{E}|S_n|^\nu M(|S_n|) \sim (n\mu)^\nu M(n\mu), \quad n \rightarrow \infty.$$

From this asymptotic relation and the assumption  $A \in \mathcal{D}(\nu, M)$  we are easily led to the conclusion

$$\sum_{n=1}^{\infty} (n\mu)^\nu M(n\mu) c_n < \infty.$$

The necessity part of the theorem then follows from the easily proved fact that  $M(n\mu) \asymp M(n)$ . It should also be obvious at this stage how to prove the sufficiency part of the theorem.

Finally we must deal with the case  $\nu = 0$  and  $\mathcal{E}X_n = \infty$ . We have, by the monotonicity of  $M(x)$ ,

$$\begin{aligned} M(|S_n|) &\leq M(|X_1| + \dots + |X_n|) \\ &= \{|X_1| + \dots + |X_n|\} \left\{ \frac{M(|X_1| + \dots + |X_n|)}{|X_1| + \dots + |X_n|} \right\}. \end{aligned}$$



But, since  $M(x)/x$  may be supposed non-increasing for all  $x \geq 0$ , we have inequalities like

$$|X_1| \left\{ \frac{M(|X_1| + \dots + |X_n|)}{|X_1| + \dots + |X_n|} \right\} \leq |X_1| \left\{ \frac{M(|X_1|)}{|X_1|} \right\}$$

and thus discover that

$$M(|S_n|) \leq M(|X_1|) + \dots + M(|X_n|).$$

Hence, if  $\mathcal{E}M(|X_1|) < \infty$ , it follows that  $\mathcal{E}M(|S_n|) = O(n)$  and the theorem is complete.

5. Some results on random walks; proof of Theorem 1.

We continue the assumption that  $\{X_n\}$  is a sequence of independent and identically distributed random variables; we write  $F(x) = P\{X_1 \leq x\}$  for their common d.f. and  $S_n = X_1 + \dots + X_n$ , as usual, for the partial sums. We suppose  $0 < \mathcal{E}X_1 \leq +\infty$ .

The following Theorem A is an easy consequence of Theorem 6 of Smith (1967). In that paper it is supposed that  $\mathcal{E}X_1$  is finite whereas we are here making no such assumption. If  $\mathcal{E}X_1 = +\infty$ , however, a suitable truncation argument suggests itself readily. It will be interesting to see how much is a consequence of this Theorem A.

Theorem A Let  $\phi(x) \in \mathcal{M}^*$  and suppose

$$\int_{-\infty}^0 |x| \phi(|x|) dF(x) < \infty.$$

If, for any constant  $\gamma < \mathcal{E}X_1$ , we write

$$\Lambda(x) = \sum_{n=1}^{\infty} \frac{1}{n} P\{S_n - n\gamma \leq x\},$$

then  $\Lambda(0) < \infty$  and

$$\int_{-\infty}^0 \phi(|x|) d\Lambda(x) < \infty.$$

If  $\mathcal{E}X_1 = \infty$  we may take  $\gamma$  arbitrarily large.

We note that  $\Lambda(x)$  is obviously a non-decreasing function of  $x \leq 0$  and that  $\phi(|x|) \Lambda(x) \rightarrow 0$  as  $x \rightarrow -\infty$ . Thus, after an integration by parts, the conclusion of Theorem A can be rewritten:

$$\int_{-\infty}^0 \Lambda(x) |d\phi(|x|)| < \infty.$$

This means

$$\sum_{n=1}^{\infty} \frac{1}{n} \int_{-\infty}^0 P\{S_n - n\gamma \leq x\} |d\phi(|x|)| < \infty$$

so that

$$\sum_{n=1}^{\infty} \frac{1}{n} \int_{-n\gamma}^0 P\{S_n - n\gamma \leq x\} |d\phi(|x|)| < \infty$$

and a fortiori

$$(5.1) \quad \sum_{n=1}^{\infty} \frac{\phi(n\gamma) - \phi(0)}{n} P\{S_n \leq 0\} < \infty.$$

Let  $M_{\nu}(x)$  represent  $x^{\nu} M(x)$  in  $\mathcal{M}^*$ , where  $M(x)$  is as defined in the enunciation of Theorem 1. Then we deduce from Theorem A and (5.1) the following:

Result 1 (1.1) implies (1.2).

Elsewhere in Smith (1967) it is shown that the assumption  $E X_1 > 0$  implies the existence of constants  $\rho > 0$ ,  $\eta > 0$ , such that

$$P\{S_n \leq 0\} \geq \rho n P\{X_1 \leq -n\eta\}.$$

Thus, if (1.2) holds, we must have

$$\sum_{n=1}^{\infty} M_{\nu}(2n\eta) P\{X_1 \leq -n\eta\} < \infty,$$

since  $M_{\nu}(2n) \asymp M_{\nu}(n)$  as  $n \rightarrow \infty$ . Therefore

$$\sum_{n=1}^{\infty} \left\{ \sum_{r=1}^n M_{\nu}(2rn) \right\} P\{-(n+1)n < X_1 \leq -n\} < \infty.$$

But, for large  $n$ ,

$$\sum_{r=1}^n M_{\nu}(2rn) \geq \frac{1}{2}(n-1)nM_{\nu}(n).$$

Hence

$$\sum_{n=1}^{\infty} nM_{\nu}(n)P\{-(n+1)n < X_1 \leq -n\} < \infty.$$

We have therefore established:

Result 2 (1.2) implies (1.1).

Let us now define, for  $n = 1, 2, \dots$ ,

$$T_n = \max\{S_1, S_2, \dots, S_n\}$$

$$\pi_n = P\{T_n \leq 0\}$$

$$\lambda_n = \frac{1}{n} P\{S_n \leq 0\}.$$

Then a familiar identity of Spitzer (1956) states that, for  $0 \leq s < 1$ ,

$$(5.2) \quad \sum_{n=1}^{\infty} \pi_n s^n = \exp \sum_{n=1}^{\infty} \lambda_n s^n.$$

From this identity, on comparing coefficients, it is clear that  $\pi_n \geq \lambda_n$  for all  $n$ . Thus we have:

Result 3 (1.3) implies (1.2)

Finally, suppose again that (1.2) holds. Then we can set

$$\Lambda = \sum_{n=1}^{\infty} \lambda_n < \infty$$

and

$$\beta(s) = \sum_{n=1}^{\infty} (\lambda_n \Lambda^{-1}) s^n, \quad 0 \leq s \leq 1,$$

will be a probability generating function in the class  $\mathcal{G}(\nu, M)$ . Furthermore

$$\gamma(s) = e^{\Lambda(s-1)}$$

is the familiar Poisson probability generating function and  $\gamma(s)$  is plainly a member of  $\mathcal{G}(\nu, M)$ . From Corollary 3.1 we can infer that  $\gamma(\beta(s))$  is another member of  $\mathcal{G}(\nu, M)$  if only we can be sure that the "mean" of  $\beta(s)$  is finite, i.e. that

$$\sum_{n=1}^{\infty} n \lambda_n < \infty.$$

This is the same as the statement

$$(5.3) \quad \sum_{n=1}^{\infty} P\{S_n \leq 0\} < \infty.$$

When  $\nu \geq 1$ , (5.3) is a consequence of (1.2), which we are assuming true.

Thus, when  $\nu \geq 1$ , we have

Result 4 (1.2) implies (1.3)

However, we must also prove Result 4 when  $\nu = 0$ . In this case (5.3) is, in general, false. However, if we expand the Poisson generating function  $\gamma(s) = \sum_0^{\infty} c_n s^n$  it is obviously true that  $\sum_1^{\infty} n c_n < \infty$ . Thus the case  $\nu = 0$  of Corollary 3.1 completes the proof of Result 4.

Theorem 1 follows from Results 1, 2, 3, and 4.

## References

- W. L. Smith (1967), A theorem on functions of characteristic functions and its application to some renewal theoretic random walk problems, Proceedings of the Fifth Berkeley Symposium on Mathematical Statistics and Probability, Vol. II, Part 2, pp. 265-309, Berkeley, University of California Press.
- W. L. Smith and W. E. Wilkinson (1965), Proceedings of the Symposium on Congestion Theory, Chapel Hill, University of North Carolina Press.
- F. Spitzer (1956), A combinatorial lemma and its application to probability theory, Trans. Amer. Math. Soc., 82, 323-339.

**UNCLASSIFIED**

Security Classification

**DOCUMENT CONTROL DATA - R & D**

*(Security classification of title, body of abstract and indexing annotation must be entered when the overall report is classified)*

1. ORIGINATING ACTIVITY (Corporate author)

2a. REPORT SECURITY CLASSIFICATION

**UNCLASSIFIED**

2b. GROUP

3. REPORT TITLE

**SOME RESULTS USING GENERAL MOMENT FUNCTIONS**

4. DESCRIPTIVE NOTES (Type of report and inclusive dates)

**Technical Report**

5. AUTHOR(S) (First name, middle initial, last name)

**Smith, Walter Laws**

6. REPORT DATE

**November 1967**

7a. TOTAL NO. OF PAGES

7b. NO. OF REFS

8a. CONTRACT OR GRANT NO.

**NONR-855-(09)**

b. PROJECT NO. **RR-003-05-01**

9a. ORIGINATOR'S REPORT NUMBER(S)

**Institute of Statistics Mimeo Series  
Number 554**

9b. OTHER REPORT NO(S) (Any other numbers that may be assigned this report)

10. DISTRIBUTION STATEMENT

**Distribution of the document is unlimited**

11. SUPPLEMENTARY NOTES

12. SPONSORING MILITARY ACTIVITY

**Logistics and Mathematical Statistics  
Branch Office of Naval Research  
Washington, D. C. 20360**

13. ABSTRACT

Let  $\{X_n\}$  be a sequence of independent and identically distributed random variables such that  $0 < \mu = \mathcal{E}X_n \leq +\infty$  and write  $S_n = X_1 + X_2 + \dots + X_n$ . Let  $\nu \geq 0$  be an integer and let  $M(x)$  be a non-decreasing function of  $x \geq 0$  such that  $M(x)/x$  is non-increasing and  $M(0) > 0$ . Then if  $\mathcal{E}|X_1|^\nu M(|X_1|) < \infty$  and  $\mu < \infty$  it follows that  $\mathcal{E}|S_n|^\nu M(|S_n|) \sim (n\mu)^\nu M(n\mu)$  as  $n \rightarrow \infty$ . If  $\mu = \infty$  ( $\nu = 0$ ) then  $\mathcal{E}M(|S_n|) = o(n)$ . A variety of results stem from this main theorem (Theorem 2), concerning a closure property of probability generating functions and a random walk result (Theorem 1) connected with queues.

14. KEY WORDS	LINK A		LINK B		LINK C	
	ROLE	WT	ROLE	WT	ROLE	WT
Renewal Theory						
Random Walks						
Moments						
Probability Generating Functions						